# Rijksuniversiteit Groningen Statistiek 

Hertentamen

## RULES FOR THE EXAM:

- The use of a normal, non-graphical calculator is permitted.
- This is a CLOSED-BOOK exam.
- At the end of the exam you can find a chi-squared table.
- Your exam mark : $10+90 \times$ your score/75.


## 1. Asymptotic distribution Likelihood Ratio Test statistic 20 Marks.

Let $X=\left(X_{1}, \ldots, X_{n}\right)$ be the observed data, such that

$$
X_{i} \stackrel{\text { i.i.d. }}{\sim} f_{\theta_{0}}
$$

Let $f_{\theta}$ be twice continuously differentiable with support not depending on $\theta$.
Let $\hat{\theta}=\hat{\theta}(X)$ be the maximum likelihood estimate of $\theta$.
(a) Derive the second order Taylor expansion of the $\log$-likelihood $\ell(\theta)=\log f_{\theta}\left(X_{1}, \ldots, X_{n}\right)$ at the true value $\ell\left(\theta_{0}\right)$ around the MLE $\hat{\theta}$. [5 Marks]
(b) Show that $\frac{1}{n} \frac{d^{2} \ell}{d \theta^{2}}\left(\theta_{0}\right) \rightarrow-I\left(\theta_{0}\right)=\left.E \frac{d^{2}}{d \theta^{2}} \log f_{\theta}\left(X_{1}\right)\right|_{\theta=\theta_{0}}$ as $n \rightarrow \infty$. [5 Marks]
(c) Use the second order Taylor expansion to show that

$$
-2 \log L R T \approx-\left(\hat{\theta}-\theta_{0}\right)^{2} \frac{d^{2} \ell}{d \theta^{2}}(\hat{\theta})
$$

where LRT is the likelihood ratio test statistic.[5 Marks]
(d) Taking the approximation in (c) as an equality, use (b) and (c) together with the asymptotic efficiency of the MLE $\hat{\theta}$ to show that

$$
-2 \log L R T \rightarrow \chi_{1}^{2}
$$

in distribution as $n \rightarrow \infty$. [5 Marks]
2. Point estimation 10 Marks.

Let $X_{1}, \ldots, X_{n}$ be a sample of independent, identically distributed random variables, with $\operatorname{Unif}(0, \theta)$ density.
(a) Determine the Method of Moments estimator $\hat{\theta}$ of $\theta$. [5 Marks]
(b) Determine whether $\hat{\theta}$ is consistent. [5 Marks]

## 3. Survival regression 30 Marks.

Let $\left(Y_{1}, x_{1}\right), \ldots,\left(Y_{n}, x_{n}\right)$ be the data, where $\left\{Y_{i}\right\}_{i=1}^{n}$ are independently and exponentially distributed random variables in the following way:

$$
Y_{i} \sim E x\left(\lambda x_{i}\right), \quad i=1,2, \ldots, n
$$

i.e.

$$
f_{Y_{i}}(y)=\lambda x_{i} e^{-\lambda x_{i} y} 1_{y \geq 0} .
$$

The known constants $\left\{x_{i}\right\}$ are strictly positive.
(a) Derive the maximum likelihood estimator for $\lambda$. [Hint: Don't forget to show that this is really a maximum.] [10 Marks]
(b) Determine whether the MLE is a sufficient statistic. [5 Marks]
(c) Exponential distributions are often used to model survival times. A researcher wants to find the relationship between the price of a light bulb and its life span. She samples 40 different light bulbs and hypothesizes the above model, whereby for $i=1, \ldots, 40$

- $x_{i}=\frac{1}{\text { (euro) price of light bulb } i}$
- $Y_{i}=$ life span of light bulb $i$ (in years).

The plot of the relationship between price and lifespan is given in the plot below. We are given that $\sum_{i=1}^{40} x_{i} y_{i}=80$.

## Life span vs price of light bulb



Figure 1: 40 measurements of light bulb life spans versus their price.
i. A consumer organization wants to know whether, based on these data, it can reject $\lambda=1$. Based on the asymptotic distribution of the MLE as test-statistic, set up a hypothesis test with null and alternative hypothesis to see if there is
sufficient evidence for a relationship at a significance level of $\alpha=0.05$ (hint: the standard normal 0.975 quantile is 1.96). [10 Marks]
ii. Calculate the (numeric!) $95 \%$ confidence interval for $\lambda$ based on the asymptotic distribution of the likelihood ratio statistic. Draw this confidence interval in a plot of the likelihood ratio statistic versus $\lambda$. [Hint: you are allowed to "read off" the numeric CI from this plot, but provide the inequality that you would like to solve.] [5 Marks]
4. Optimal testing 15 Marks. Consider a single observation $X$ from a exponential distribution with mean $\mu$, i.e. with density

$$
f_{X}(x)=\frac{1}{\mu} e^{-x / \mu}, \quad x>0
$$

We want to test the following hypotheses:

$$
\begin{array}{ll}
H_{0}: & \mu=1 \\
H_{1}: & \mu=2
\end{array}
$$

(a) We want to perform an optimal test with a significance level of at most $5 \%$ of the null hypothesis against the alternative. Determine the critical region. [10 Marks]
(b) What is the power of this test? [5 Marks]

Below statistical tables which may be used in the calculations.

| $\nu \backslash \alpha$ | 0.995 | 0.99 | 0.975 | 0.95 | 0.05 | 0.025 | 0.01 | 0.005 |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| 1 | 0.000 | 0.000 | 0.001 | 0.004 | 3.841 | 5.024 | 6.635 | 7.879 |
| 2 | 0.010 | 0.020 | 0.051 | 0.103 | 5.991 | 7.378 | 9.210 | 10.597 |
| 3 | 0.072 | 0.115 | 0.216 | 0.352 | 7.815 | 9.348 | 11.345 | 12.838 |
| 4 | 0.207 | 0.297 | 0.484 | 0.711 | 9.488 | 11.143 | 13.277 | 14.860 |
| 5 | 0.412 | 0.554 | 0.831 | 1.145 | 11.070 | 12.833 | 15.086 | 16.750 |
| 10 | 2.156 | 2.558 | 3.247 | 3.940 | 18.307 | 20.483 | 23.209 | 25.188 |

Table 1: Values of $\chi_{\alpha, \nu}^{2}$ as found in the book: the entries in the table correspond to values of $x$, such that $P\left(\chi_{\nu}^{2}>x\right)=\alpha$, where $\chi_{\nu}^{2}$ correspond to a chi-squared distributed variable with $\nu$ degrees of freedom.

